Chapter 8 - Lecture 2 (2) Large-Sample Confidence Intervals for a Population Proportion

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1 Review: general population with σ^2 unknown

2 CI for proportion: traditional interval

3 CI for proportion: score interval

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Review

Without normality assumption, when *n* is large and σ^2 is unknown, we still get approximate z CI for the population mean μ .

An approximate $100(1 - \alpha)$ % confidence interval for μ is:

$$(\bar{x} - z_{\alpha/2}s/\sqrt{n}, \bar{x} + z_{\alpha/2}s/\sqrt{n})$$

The above CI is valid when CLT is valid.

For estimating proportion, each random variable is bernoulli variable, with 1, if event happens and 0 otherwise. The data we collect will be a sequence of 0 and 1. The estimator is

$$\hat{p} = \bar{X} = rac{\mathsf{Number of nonzero}}{\mathsf{Sample size}}$$

For the proportion, we know that to apply CLT, we have to check np > 10, n(1-p) > 10. And then we plug in the $s = \sqrt{\hat{p}(1-\hat{p})}$ to get the traditional interval.

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The **traditional way** of approximate $100(1 - \alpha)$ % confidence interval for *p* is:

$$P\left(-z_{\alpha/2} < \frac{\bar{X} - \mu}{S/\sqrt{n}} < z_{\alpha/2}\right) \approx 1 - \alpha$$

$$\rightarrow \quad (\hat{p} - z_{\alpha/2}s/\sqrt{n}, \hat{p} - z_{\alpha/2}s/\sqrt{n})$$

$$\rightarrow \quad (\hat{p} - z_{\alpha/2}\sqrt{\hat{p}(1-\hat{p})}/\sqrt{n}, \hat{p} - z_{\alpha/2}\sqrt{\hat{p}(1-\hat{p})}/\sqrt{n})$$

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To get the score interval (default in book),

$$P\left(-z_{\alpha/2} < \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} < z_{\alpha/2}\right) \approx 1 - \alpha$$

$$\rightarrow P\left(-z_{\alpha/2} < \frac{\hat{p} - p}{\sqrt{p(1-p)}/\sqrt{n}} < z_{\alpha/2}\right) \approx 1 - \alpha$$

Resolve p will get a (score) interval for a population proportion p with a confidence level approximately $(1 - \alpha)$ Cl is :

$$rac{\hat{
ho}+rac{z_{lpha/2}^2}{2n}}{1+rac{z_{lpha/2}^2}{n}}\pm z_{lpha/2}rac{\sqrt{rac{\hat{
ho}(1-\hat{
ho})}{n}}+rac{z_{lpha/2}^2}{4n^2}}{1+rac{z_{lpha/2}^2}{n}}$$

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Example 8.8

The article "Repeatability and Reproducibility for Pass/Fail Data" (*J. Testing Eval.*, 1997: 151–153) reported that in n = 48 trials in a particular laboratory, 16 resulted in ignition of a particular type of substrate by a lighted cigarette. Let p denote the long-run proportion of all such trials that would result in ignition. A point estimate for p is $\hat{p} = 16/48 = .333$. A confidence interval for p with a confidence level of approximately 95% is

$$\frac{.333 + 1.96^2/96}{1 + 1.96^2/48} \pm 1.96 \frac{\sqrt{(.333)(.667)/48 + 1.96^2/(4 \cdot 48^2)}}{1 + 1.96^2/48}$$
$$= .346 \pm .129 = (.217, .475)$$

The traditional interval is

$$.333 \pm 1.96\sqrt{(.333)(.667)/48} = .333 \pm .133 = (.200, .466)$$

These two intervals would be in much closer agreement were the sample size substantially larger.

 Figure: Modern Mathematical Statistics with Applications, 2rd, P397a

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